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Tracking with Non-linear Dynamic Models

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Tracking with Non-linear Dynamic Models

- Distribution propagation
- Problems with non-linearities
- Sampling densities
- Particle filtering
- Tracking people

Distribution propogation



[Isard 1998]

Distribution propogation



[Isard 1998]

A little nonlinearity becomes very non-Gaussian



 $x_{i+1} = x_i + 0.1 \sin x_i$

time evolution of the state of a set of 100 points



EKF

Linearize system at each time point to form an Extended Kalman Filter (EKF)

– Compute Jacobian matrix

 $\mathcal{J}(\boldsymbol{g}; \boldsymbol{x}_j)$ whose (l,m)'th value is $\frac{\partial f_l}{\partial x_m}$ evaluated at \boldsymbol{x}_j – use this for forward model at each step in KF

Useful in many engineering applications, but not as successful in computer vision....

Representing non-linear Distributions



Representing non-linear Distributions

Unimodal parametric models fail to capture realworld densities...



Representing non-linear Distributions

Mixture models are appealing, but very hard to propagate analytically!



[but see Cham and Rehg's MHT approach]

Representing Distributions using Weighted Samples

Rather than a parametric form, use a set of samples to represent a density:



Representing Distributions using Weighted Samples

Rather than a parametric form, use a set of samples to represent a density:



1. a set of sample locations from what distribution?

2. weights associated with those locations how does that distribution effect the weights?





Expectation over sampled density

If
$$p_f(\mathbf{X}) = \frac{f(\mathbf{X})}{\int f(\mathbf{U})d\mathbf{U}}$$

$$w^i = f(\boldsymbol{u}^i) / s(\boldsymbol{u}^i)$$

$$\begin{split} \mathbf{E}_{p_f}\left[g\right] &= \int g(\boldsymbol{U}) p_f(\boldsymbol{U}) d\boldsymbol{U} \\ &= \frac{\int g(\boldsymbol{U}) f(\boldsymbol{U}) d\boldsymbol{U}}{\int f(\boldsymbol{U}) d\boldsymbol{U}} \\ &= \mathbf{E}\left[\frac{\sum_i g(\boldsymbol{u}_i) w_i}{\sum_i w_i}\right] \\ &\approx \frac{\sum_i g(\boldsymbol{u}_i) w_i}{\sum_i w_i} \end{split}$$

$$\begin{split} \mathbf{E}\left[\frac{1}{N}\sum_{i}w^{i}\right] &= \int 1\frac{f(\boldsymbol{U})}{s(\boldsymbol{U})}s(\boldsymbol{U})d\boldsymbol{U}\\ &= \int f(\boldsymbol{U})d\boldsymbol{U} \end{split}$$

Expectation over sampled density

We have a representation of a probability distribution

$$p_f(\boldsymbol{X}) = \frac{f(\boldsymbol{X})}{\int f(\boldsymbol{U}) d\boldsymbol{U}}$$

by a set of weighted samples

$$\left\{ \left({{{oldsymbol{u}}^{i}},{{w}^{i}}}
ight)
ight\}$$

where $\boldsymbol{u}^i \sim s(\boldsymbol{u})$ and $w^i = f(\boldsymbol{u}^i)/s(\boldsymbol{u}^i)$. Then:

$$\int g(oldsymbol{U}) p_f(oldsymbol{U}) doldsymbol{U} pprox rac{\sum_{i=1}^N g(oldsymbol{u}^i) w^i}{\sum_{i=1}^N w^i}$$

Sampled representation of a probability distribution

Represent a probability distribution

$$p_f(\boldsymbol{X}) = \frac{f(\boldsymbol{X})}{\int f(\boldsymbol{U}) d\boldsymbol{U}}$$

by a set of N weighted samples

$$\left\{ \left({{oldsymbol{u}}^{i},{w}^{i}}
ight)
ight\}$$

where $\boldsymbol{u}^i \sim s(\boldsymbol{u})$ and $w^i = f(\boldsymbol{u}^i)/s(\boldsymbol{u}^i)$.

Marginalizing a sampled density

If we have a sampled representation of a joint density

$$\left\{((\boldsymbol{m}^i, \boldsymbol{n}^i), w^i)
ight\}$$

and we wish to marginalize over one variable:

$$p_f(\boldsymbol{M}) = \int p_f(\boldsymbol{M}, \boldsymbol{N}) d\boldsymbol{N}$$

we can simply ignore the corresponding components of the samples (!):

$$egin{aligned} &\int g(oldsymbol{M}) p_f(oldsymbol{M}) doldsymbol{M} &= \int \int g(oldsymbol{M}) p_f(oldsymbol{M},oldsymbol{N}) doldsymbol{N} doldsymbol{M} \ &= \int \int g(oldsymbol{M}) p_f(oldsymbol{M},oldsymbol{N}) doldsymbol{N} doldsymbol{M} \ &pprox rac{\sum_{i=1}^N g(oldsymbol{m}^i) w^i}{\sum_{i=1}^N w^i} \end{aligned}$$

Marginalizing a sampled density

Assume we have a sampled representation of a distribution

 $p_f(\boldsymbol{M}, \boldsymbol{N})$

given by

$$\left\{((\boldsymbol{m}^{i},\boldsymbol{n}^{i}),w^{i})\right\}$$

Then

$$\left\{ (\boldsymbol{m}^{i},w^{i})
ight\}$$

is a representation of the marginal,

$$\int p_f(oldsymbol{M},oldsymbol{N})doldsymbol{N}$$

Sampled Bayes

Transforming a Sampled Representation of a Prior into a Sampled Representation of a Posterior:

$$\int g(\boldsymbol{U})p(\boldsymbol{U}|\boldsymbol{V}=v_0)d\boldsymbol{U}=\frac{1}{K}\int g(\boldsymbol{U})p(\boldsymbol{V}=\boldsymbol{v}_0|\boldsymbol{U})p(\boldsymbol{U})d\boldsymbol{U}$$

$$\approx \frac{1}{K} \frac{\sum_{i=1}^{N} g(\boldsymbol{u}^{i}) p(\boldsymbol{V} = \boldsymbol{v}_{0} | \boldsymbol{u}^{i}) w^{i}}{\sum_{i=1}^{N} w^{i}}$$

$$pprox rac{\sum_{i=1}^N g(oldsymbol{u}^i) p(oldsymbol{V} = oldsymbol{v}_0 | oldsymbol{u}^i) w^i}{\sum_{i=1}^N p(oldsymbol{V} = oldsymbol{v}_0 | oldsymbol{u}^i) w^i}$$

Sampled Bayes

Assume we have a representation of p(U) as

$$\left\{(\boldsymbol{u}^i, w^i)\right\}$$

Assume we have an observation $V = v_0$, and a likelihood model p(V|U).

The posterior, $p(\boldsymbol{U}|\boldsymbol{V}=\boldsymbol{v}_0)$ is represented by

$$\left\{(\boldsymbol{u}^{i}, w'^{i})\right\}$$

where

$$w'^i = p(\boldsymbol{V} = \boldsymbol{v}_0 | \boldsymbol{u}^i) w^i$$

Sampled Prediction

$$P(\mathbf{x}_{i}|\mathbf{y}_{0},...,\mathbf{y}_{i-1}) = ?$$

$$p(\mathbf{X}_{i},\mathbf{X}_{i-1}|\mathbf{y}_{0},...,\mathbf{y}_{i-1}) = p(\mathbf{X}_{i}|\mathbf{X}_{i-1})p(\mathbf{X}_{i-1}|\mathbf{y}_{0},...,\mathbf{y}_{i-1})$$

$$p(\mathbf{X}_{i-1}|\mathbf{y}_{0},...,\mathbf{y}_{i-1})$$

$$\{(\mathbf{u}_{i-1}^{k},w_{i-1}^{k})\} \longrightarrow \mathbf{x}_{i} = \mathbf{f}(\mathbf{x}_{i-1}) + \xi_{i} \longrightarrow$$

$$\{((f(\mathbf{u}_{i-1}^{k}) + \xi_{i}^{l},\mathbf{u}_{i-1}^{k}),w_{i-1}^{k})\}$$

$$p(\mathbf{X}_{i},\mathbf{X}_{i-1}|\mathbf{y}_{0},...,\mathbf{y}_{i-1})$$

Drop elements to get $P(\boldsymbol{x}_i | \boldsymbol{y}_0, \dots, \boldsymbol{y}_{i-1}) \sim = \{(f(\boldsymbol{u}_{i-1}^k) + \xi_i^l, w_{i-1}^k)\}$

Sampled Correction

Prior \rightarrow posterior Reweight with

$$p(\boldsymbol{Y}_{i} = \boldsymbol{y}_{i} | \boldsymbol{X}_{i} = \boldsymbol{s}_{i}^{k,-}) w_{i}^{k,-}$$

yielding

$$\left\{ (\boldsymbol{s}_{i}^{k,-}, p(\boldsymbol{Y}_{i} = \boldsymbol{y}_{i} | \boldsymbol{X}_{i} = \boldsymbol{s}_{i}^{k,-}) w_{i}^{k,-}) \right\}$$

Naïve PF Tracking

- Start with samples from something simple (Gaussian)
- Repeat
 - Correct

$$\left\{(\bm{s}_{i}^{k,-}, p(\bm{Y}_{i}=\bm{y}_{i}|\bm{X}_{i}=\bm{s}_{i}^{k,-})w_{i}^{k,-})\right\}$$

– Predict

$$\left\{ \left(f(\boldsymbol{u}_{i-1}^k) + \xi_i^l, w_{i-1}^k \right) \right\}$$

Doesn't work that well

Sample impoverishment

Test with linear case:



Sample impoverishment

10 of the 100 particles:



Resample the prior

In a sampled density representation, the frequency of samples can be traded off against weight:

$$(\boldsymbol{s}_k, w_k) \longrightarrow (\begin{array}{c} (\boldsymbol{s}_k, 1) \\ (\boldsymbol{s}_k, 1) \\ (\boldsymbol{s}_k, 1) \end{array} N_k \text{ copies s.t. } \frac{N_k}{\sum_k N_k} = w_k$$

 \vdots

These new samples are a representation of the same density.

I.e., make N draws with replacement from the original set of samples, using the weights as the probability of drawing a sample.

Resampling concentrates samples



A practical particle filter with resampling

Initialization: Represent $P(X_0)$ by a set of N samples

$$\left\{(s_0^{k,-}, w_0^{k,-})\right\}$$

where

$$s_0^{k,-} \sim P_s(S)$$
 and $w_0^{k,-} = P(s_0^{k,-})/P_s(S = s_0^{k,-})$

Ideally, $P(X_0)$ has a simple form and $s_0^{k,-} \sim P(X_0)$ and $w_0^{k,-} = 1$. **Prediction:** Represent $P(X_i|y_0, y_{i-1})$ by

$$\left\{(s_i^{k,-},w_i^{k,-})\right\}$$

where

$$s_i^{k,-} = f(s_{i-1}^{k,+}) + \xi_i^k$$
 and $w_i^{k,-} = w_{i-1}^{k,+}$ and $\xi_i^k \sim N(0, \Sigma_{d_i})$

Correction: Represent $P(X_i|y_0, y_i)$ by

$$\left\{(s_i^{k,+},w_i^{k,+})\right\}$$

where

$$s_i^{k,+} = s_i^{k,-}$$
 and $w_i^{k,+} = P(\boldsymbol{Y}_i = \boldsymbol{y}_i | \boldsymbol{X}_i = s_i^{k,-}) w_i^{k,-}$

Resampling: Normalise the weights so that $\sum_i w_i^{k,+} = 1$ and compute the variance of the normalised weights. If this variance exceeds some threshold, then construct a new set of samples by drawing, with replacement, N samples from the old set, using the weights as the probability that a sample will be drawn. The weight of each sample is now 1/N.

A variant



[Isard 1998]

A variant (animation)





Applications

Tracking

- hands
- bodies
- leaves

Contour tracking





Head tracking







(a)

[Isard 1998]

Leaf tracking





Hand tracking





Mixed state tracking





A drawing interface



[Isard 1998]

Articulated tracking



Interesting Extensions

- Multiple people / objects
 - state has to model multi-body configuration
 - resampling is tricky!
- Multiple modalities
 - fuse observation likelihoods
 - audio/visual localization and source separation

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[Figures from F&P except as noted]