



# Machine learning: lecture 13

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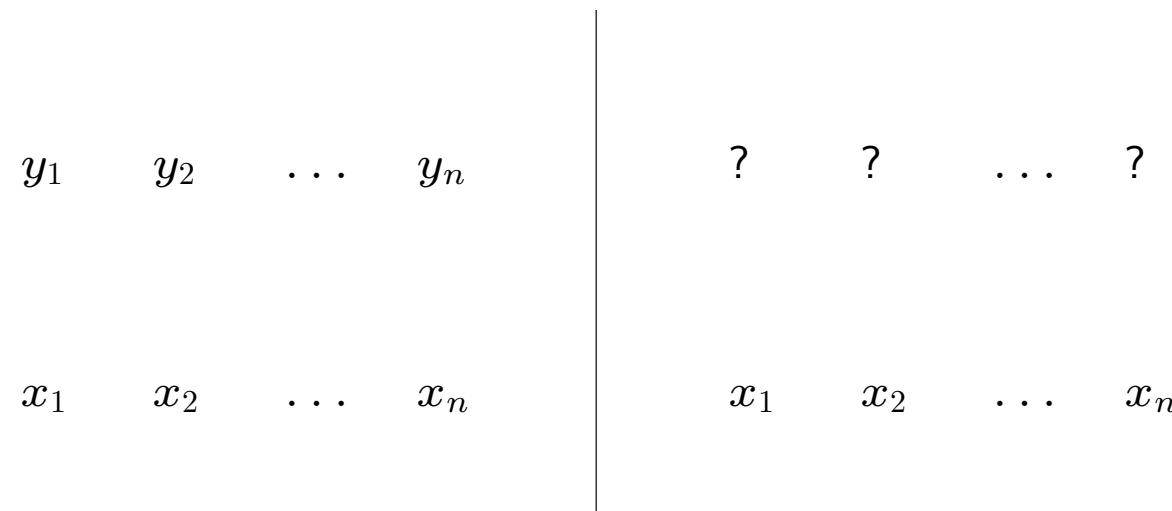
# Topics

- Complexity, compression, and model selection
  - description length
  - minimum description length principle
- Probabilistic modeling
  - mixture models, EM algorithm



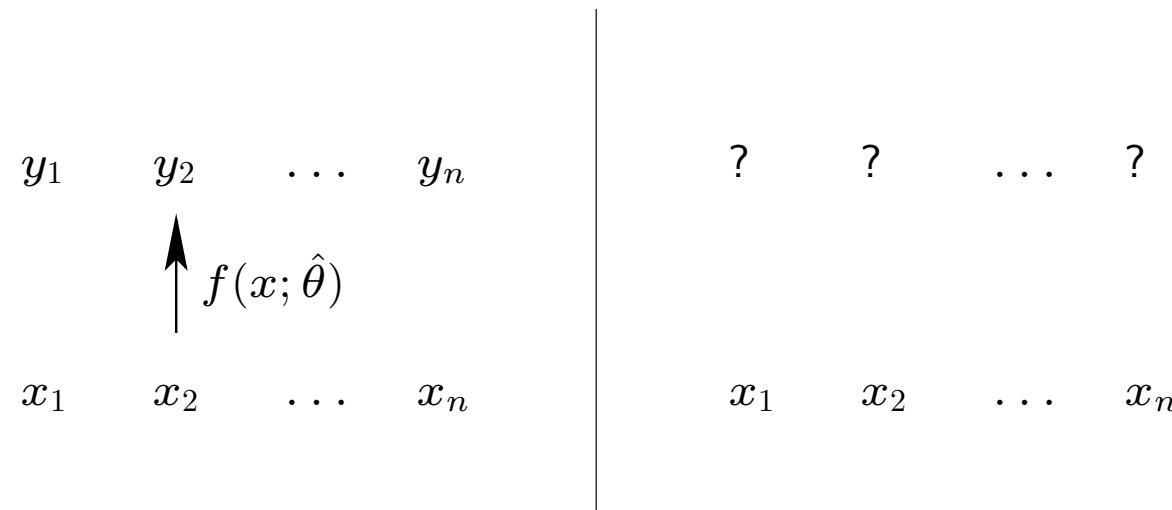
# Data compression and model selection

- We can alternatively view model selection as a problem of finding the best way of communicating the available data



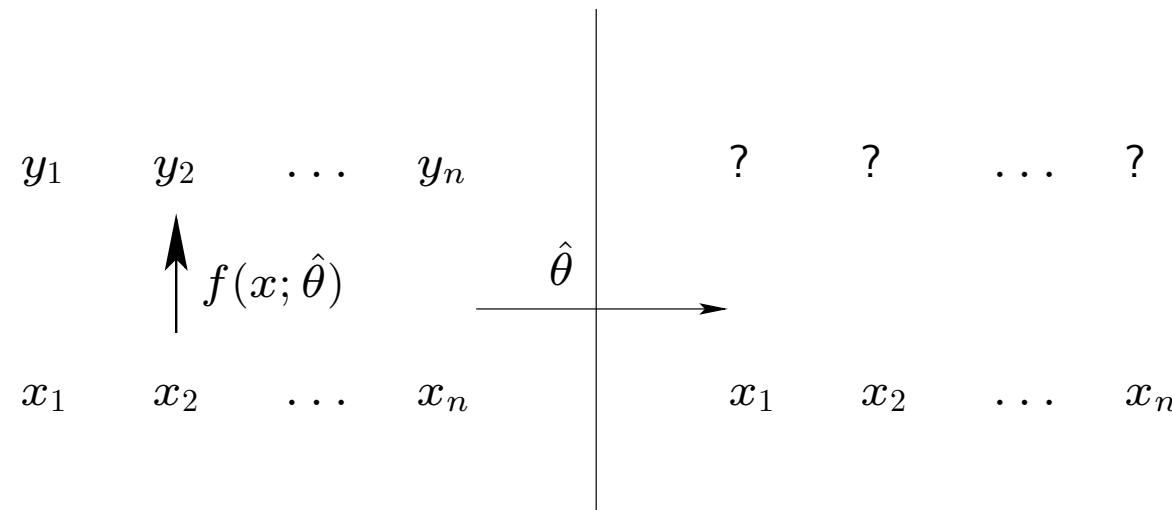
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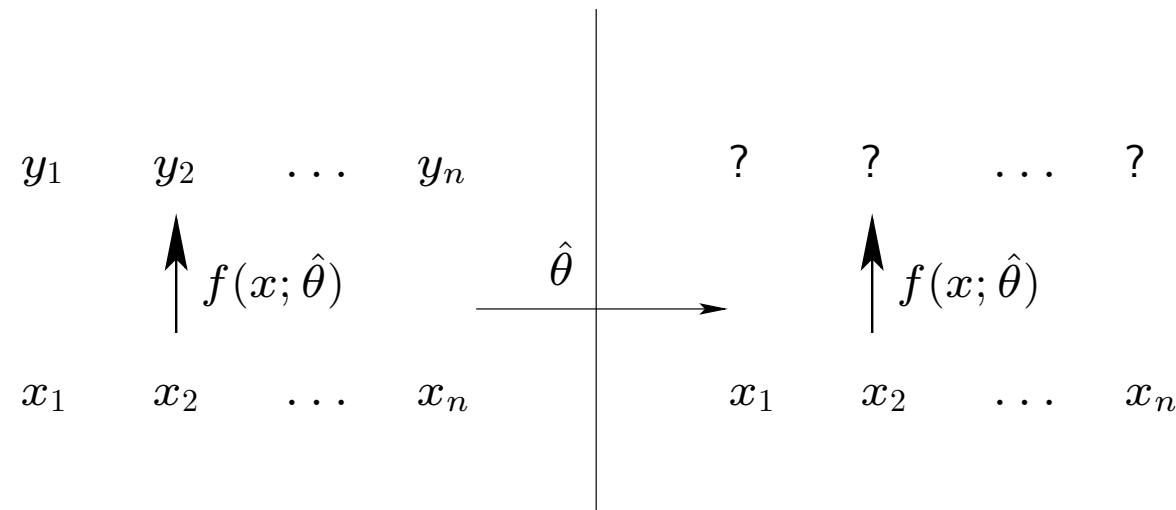
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- We can alternatively view model selection as a problem of finding the best way of communicating the available data



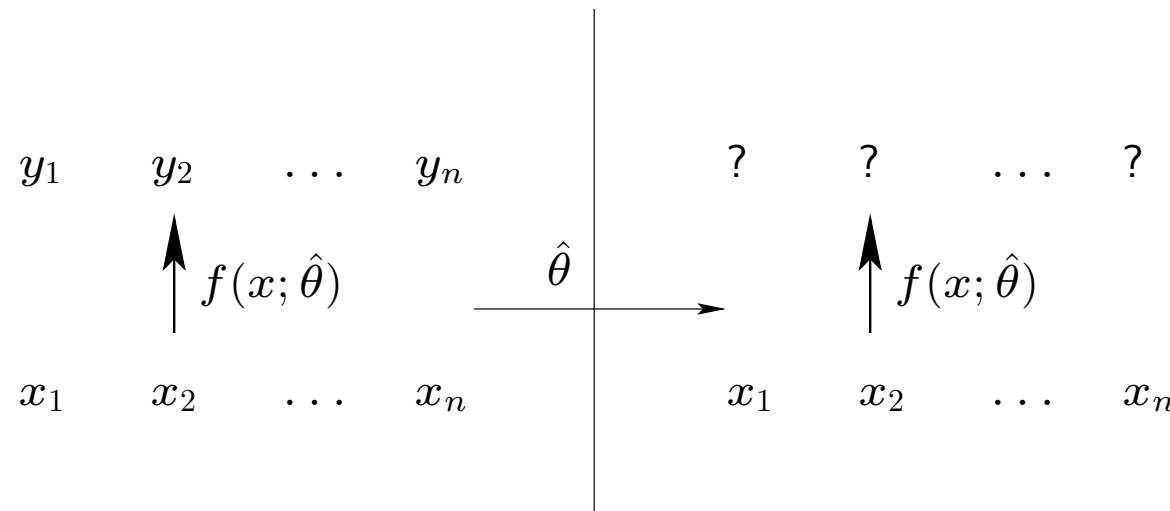
# Data compression and model selection

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# Data compression and model selection

- We can alternatively view model selection as a problem of finding the best way of communicating the available data



What is shared between the sender and the receiver?

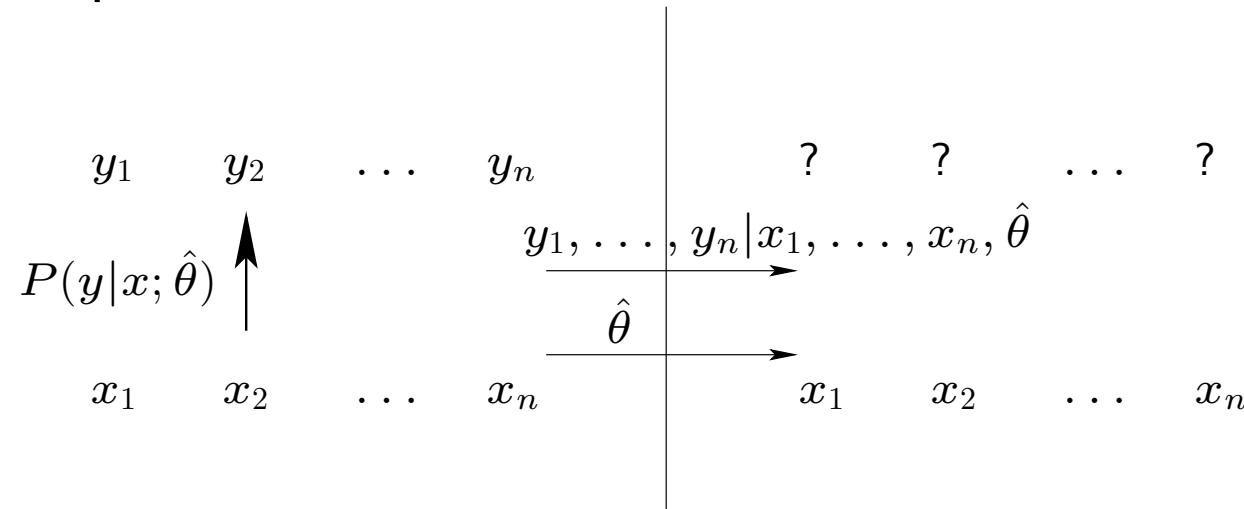
- input examples  $\mathbf{x}_1, \dots, \mathbf{x}_n$
- knowledge of function classes

What needs to be communicated?

- anything pertaining to the labels  $y_1, \dots, y_n$

# Data compression and model selection

- To communicate the labels effectively we need to cast the problem in probabilistic terms

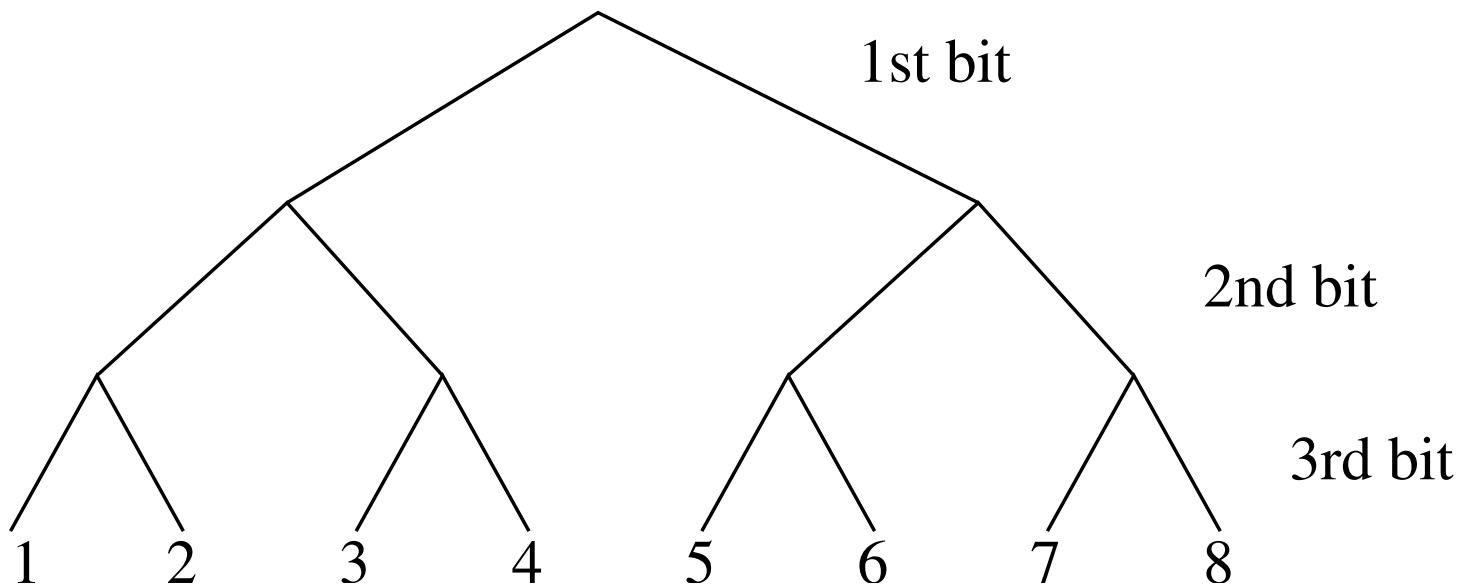


- The communication cost in bits depends on how well the model can predict the data as well as how hard it is to describe the model itself (complexity)

Total # of bits = bits to describe the data given the model  
+ bits to describe the model

# Bits and probabilities

- How many bits do we need to communicate a specific selection out of a set of eight equally likely choices?



We need  $-\log_2 P(y) = -\log_2(1/8) = 3$  bits to describe each  $y$ .



## Description length

- How many bits do we need to describe

01111111111011100111111111111110111111

- If we assume that the bits  $\{y_i\}$  in the sequence are independent random draws from  $P$ , where  $P(y = 1) = 0.5$ , then

$$\sum_{i=1}^{40} (-\log_2 P(y_i)) = 40 \text{ bits}$$

- If we assume instead that  $P(y = 1) = 0.9$ , then

$$\sum_{i=1}^{40} (-\log_2 P(y_i)) \approx 22 \text{ bits}$$

- What we assume matters a great deal.



## Conditional description length

- We can also describe outcomes conditionally, i.e., determine the number of bits we need to specify  $y$  given  $\mathbf{x}$

$$\begin{array}{ccccccc} y_1 & y_2 & y_3 & y_4 & \dots \\ \mathbf{x}_1 & \mathbf{x}_2 & \mathbf{x}_3 & \mathbf{x}_4 & \dots \end{array}$$

Assuming the labels are generated from a conditional distribution  $P(y|\mathbf{x}, \theta)$ , we need

$$\sum_i (-\log_2 P(y_i|\mathbf{x}_i, \theta))$$

bits to describe the outcomes (labels).

- The actual number of bits may vary considerably as a function of the parameters  $\theta$ .



## Description length cont'd

- We can of course find  $\hat{\theta} \in \Theta$  (the maximum likelihood parameter estimate) that minimizes the number of bits needed to describe the labels given examples

$$\sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \hat{\theta}) \right)$$

The minimizing  $\hat{\theta}$  depends on the labels and needs to be communicated as well.



## Description length cont'd

- In addition to describing the data using  $\hat{\theta}$  with

$$\sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \hat{\theta}) \right) \text{ bits},$$

we have to communicate  $\hat{\theta}$ .

$$\text{total DL} = \text{DL of data using } \hat{\theta} + \text{DL of } \hat{\theta}$$

- The description length of the parameters  $\hat{\theta}$  depends on the model (the set of distributions we are considering)
  - the more choices we have, the more bits it takes to describe any specific selection



## How to describe the parameters

- We need to encode the parameters up to a finite precision  $\delta_k = 1/2^k$ , i.e., use  $k$  significant bits (we assume here that the precision is the same for all parameters)
- With the help of a prior density  $p(\theta)$ , it takes us roughly speaking

$$-\log_2 \left( p(\theta_{\delta_k}) \delta_k^d \right)$$

bits to describe any finite precision choice  $\theta_{\delta_k}$ . Here  $d$  is the dimensionality of the parameter vector  $\theta$ .



## How to describe the parameters cont'd

- We also need to communicate our choice of precision or  $k$  since this choice may be based on the labels.

This takes us

$$\log_2^*(k) = \log_2(k) + \log_2 \log_2(k) + \dots \text{ bits}$$

(based on a specific prior over integers).



## Description length

- The total description length – bits needed to communicate the labels given examples – is given by the minimum of

$$\sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \theta_{\delta_k}) \right) - \log_2 \left( p(\theta_{\delta_k}) \delta_k^d \right) + \log_2^*(k)$$

where the minimization is taken with respect to finite precision choices  $\theta_{\delta_k}$  as well as the number of significant bits  $k$ .



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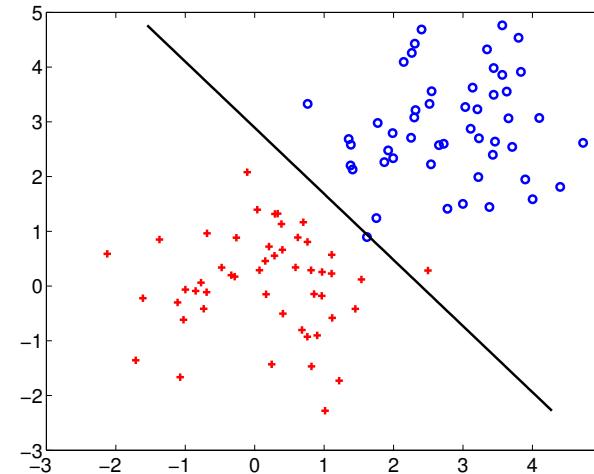
- For large  $n$  we can use the following asymptotic expansion:

$$\sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \hat{\theta}) \right) + \frac{d}{2} \log_2(n)$$

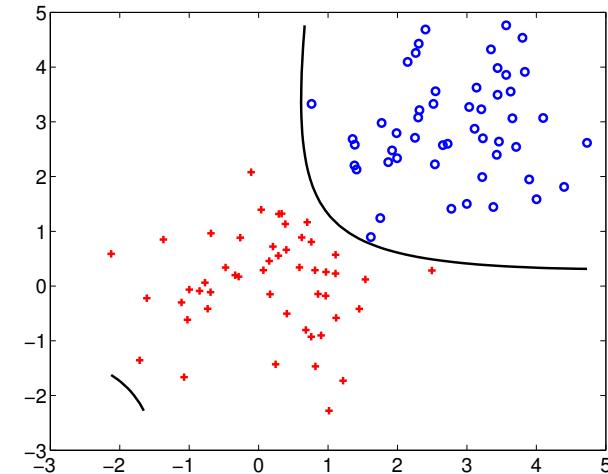
where  $\hat{\theta}$  is the maximum likelihood setting of the parameters and  $d$  is the effective number of parameters.

# Description length: example

- Example: polynomial logistic regression,  $n = 100$



linear

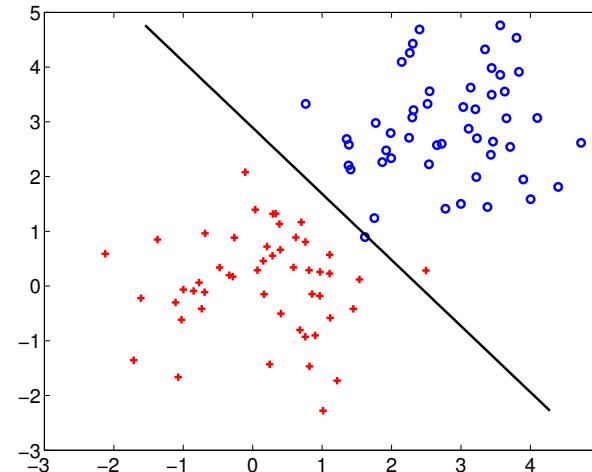


quadratic

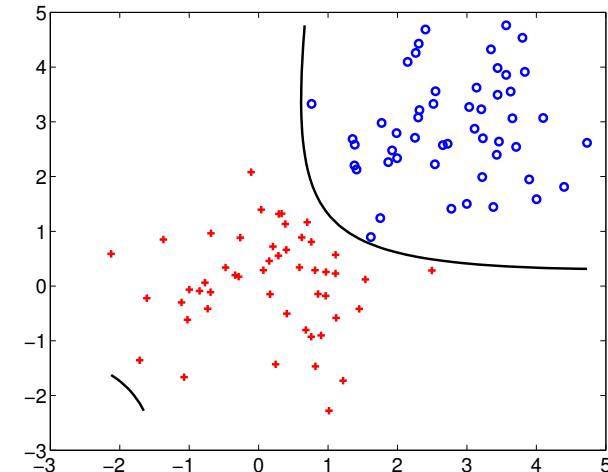
$$DL = \sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \hat{\theta}) \right) + \frac{d}{2} \log_2(n)$$

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$$DL = \sum_i \left( -\log_2 P(y_i | \mathbf{x}_i, \hat{\theta}) \right) + \frac{d}{2} \log_2(n)$$

degree	# param	DL(data)	DL(model)	MDL score
1	3	5.6 bits	9.9 bits	15.5 bits
2	6	2.4 bits	19.9 bits	22.3 bits

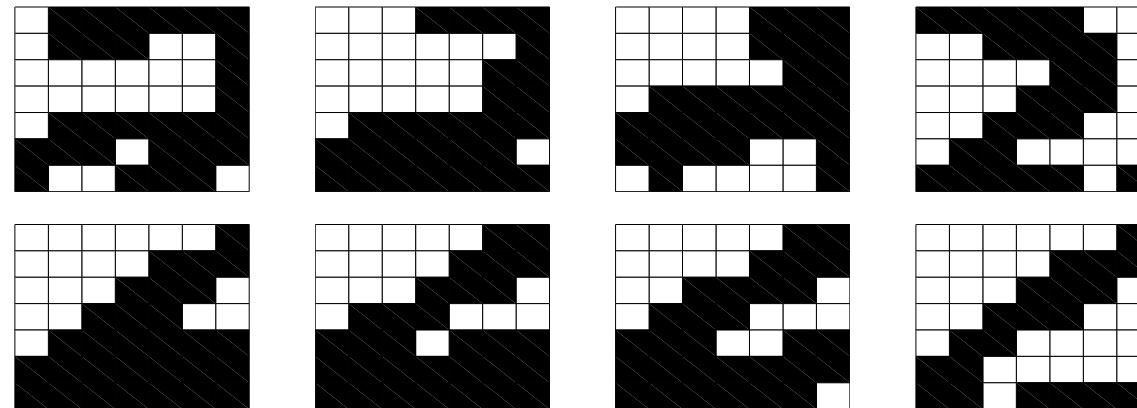


# Topics

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- Probabilistic modeling
  - mixture models, EM algorithm

# Probabilistic modeling

The digits again...



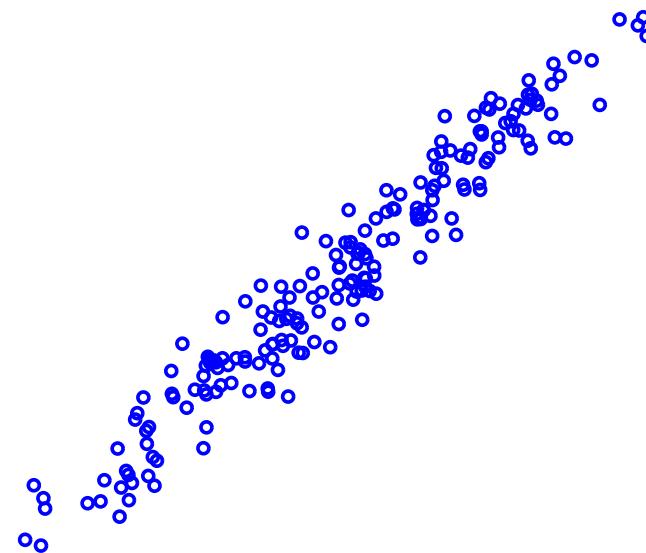
- Possible uses:
  - clustering
  - understanding the generation process of examples
  - classification via class-conditional densities
  - inference based on incomplete observations

# Parametric density models

- Probability model = a parameterized family of probability distributions
- Example: a simple multivariate Gaussian model

$$p(\mathbf{x}|\mu, \Sigma) = \frac{1}{(2\pi)^{p/2}|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu)\right\}$$

- This is a *generative model* in the sense that we can generate  $\mathbf{x}$ 's

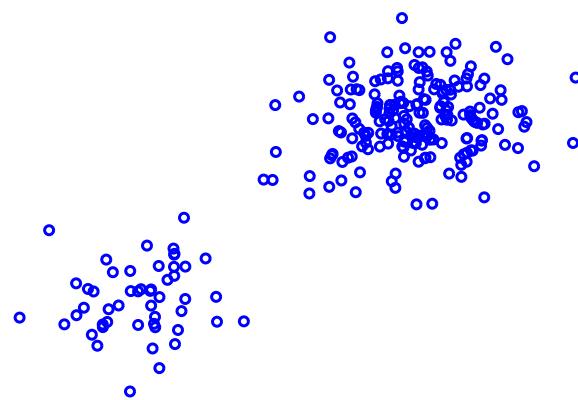


# Multi-variate density estimation

- A mixture of Gaussians model

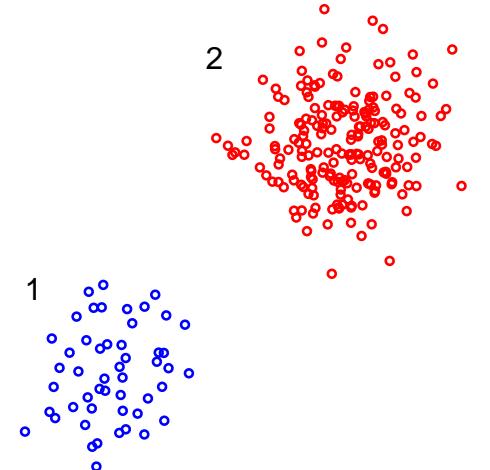
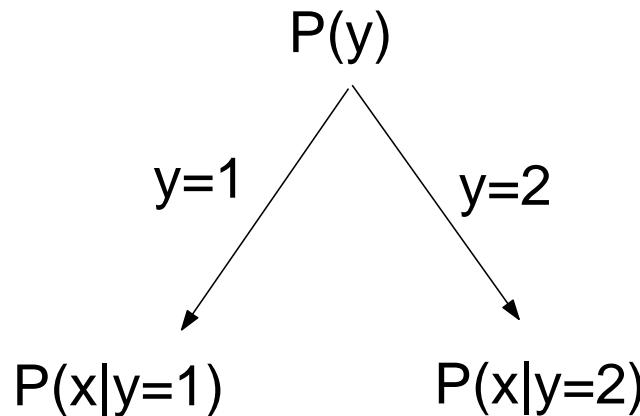
$$p(\mathbf{x}|\theta) = \sum_{i=1}^k p_j p(\mathbf{x}|\mu_j, \Sigma_j)$$

where  $\theta = \{p_1, \dots, p_k, \mu_1, \dots, \mu_k, \Sigma_1, \dots, \Sigma_k\}$  contains all the parameters of the mixture model.  $\{p_j\}$  are known as *mixing proportions or coefficients*.



# Mixture density

- Data generation process:



$$\begin{aligned}
 p(\mathbf{x}|\theta) &= \sum_{j=1,2} P(y = j) \cdot p(\mathbf{x}|y = j) \quad (\text{generic mixture}) \\
 &= \sum_{j=1,2} p_j \cdot p(\mathbf{x}|\mu_j, \Sigma_j) \quad (\text{mixture of Gaussians})
 \end{aligned}$$

- Any data point  $\mathbf{x}$  could have been generated in two ways



## Mixture density

- If we are given just  $\mathbf{x}$  we don't know which mixture component this example came from

$$p(\mathbf{x}|\theta) = \sum_{j=1,2} p_j p(\mathbf{x}|\mu_j, \Sigma_j)$$

- We can evaluate the posterior probability that an observed  $\mathbf{x}$  was generated from the first mixture component

$$\begin{aligned} P(y = 1|\mathbf{x}, \theta) &= \frac{P(y = 1) \cdot p(\mathbf{x}|y = 1)}{\sum_{j=1,2} P(y = j) \cdot p(\mathbf{x}|y = j)} \\ &= \frac{p_1 p(\mathbf{x}|\mu_1, \Sigma_1)}{\sum_{j=1,2} p_j p(\mathbf{x}|\mu_j, \Sigma_j)} \end{aligned}$$

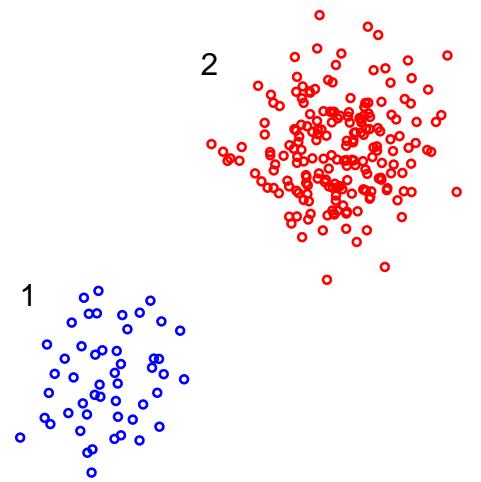
- This solves a *credit assignment* problem

# Mixture density estimation

- Suppose we want to estimate a two component mixture of Gaussians model.

$$p(\mathbf{x}|\theta) = p_1 p(\mathbf{x}|\mu_1, \Sigma_1) + p_2 p(\mathbf{x}|\mu_2, \Sigma_2)$$

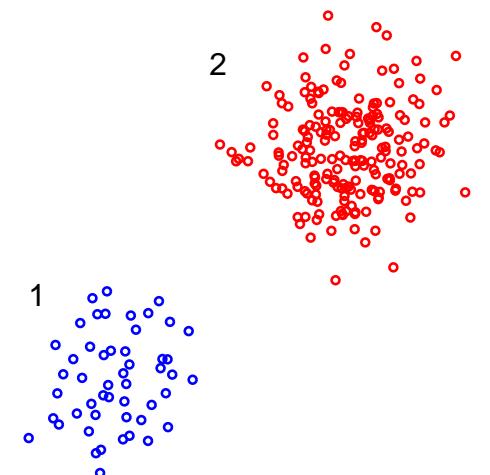
- If each example  $\mathbf{x}_i$  in the training set were labeled  $y_i = 1, 2$  according to which mixture component (1 or 2) had generated it, then the estimation would be easy.



- Labeled examples  $\Rightarrow$  no credit assignment problem

# Mixture density estimation

- When examples are labeled, we can estimate each Gaussian independently
- Let  $\delta(j|i)$  be an indicator function of whether example  $i$  is labeled  $j$ . Then for each  $j = 1, 2$



$$\hat{p}_j \leftarrow \frac{\hat{n}_j}{n}, \text{ where } \hat{n}_j = \sum_{i=1}^n \delta(j|i)$$

$$\hat{\mu}_j \leftarrow \frac{1}{\hat{n}_j} \sum_{i=1}^n \delta(j|i) \mathbf{x}_i$$

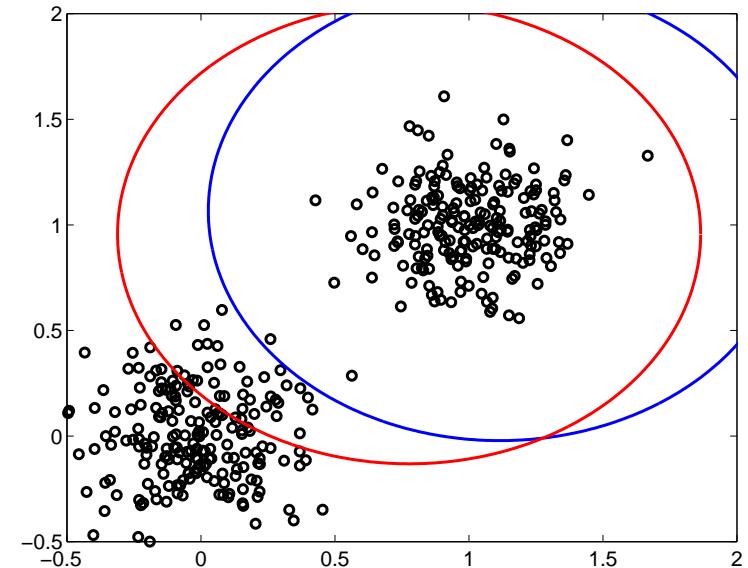
$$\hat{\Sigma}_j \leftarrow \frac{1}{\hat{n}_j} \sum_{i=1}^n \delta(j|i) (\mathbf{x}_i - \hat{\mu}_j)(\mathbf{x}_i - \hat{\mu}_j)^T$$

# Mixture density estimation: credit assignment

- Of course we don't have such labels ... but we can guess what the labels might be based on our current mixture distribution
- We get soft labels or posterior probabilities of which Gaussian generated which example:

$$\hat{p}(j|i) \leftarrow P(y_i = j | \mathbf{x}_i, \theta)$$

where  $\sum_{j=1,2} \hat{p}(j|i) = 1$  for all  $i = 1, \dots, n$ .

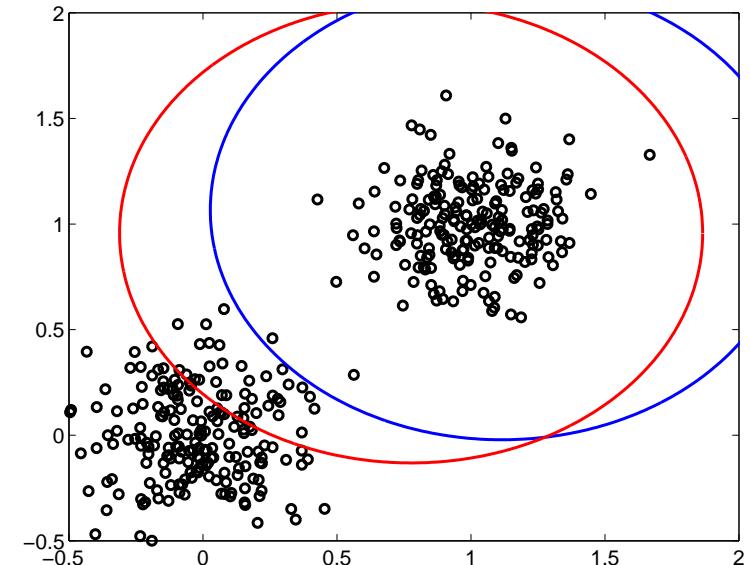


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where  $\sum_{j=1,2} \hat{p}(j|i) = 1$  for all  $i = 1, \dots, n$ .



- When the Gaussians are almost identical (as in the figure),  $\hat{p}(1|i) \approx \hat{p}(2|i)$  for almost any available point  $\mathbf{x}_i$ . Even slight differences can help us determine how we should modify the Gaussians.



# The EM algorithm

**E-step:** softly assign examples to mixture components

$$\hat{p}(j|i) \leftarrow P(y_i = j | \mathbf{x}_i, \theta), \text{ for all } j = 1, 2 \text{ and } i = 1, \dots, n$$

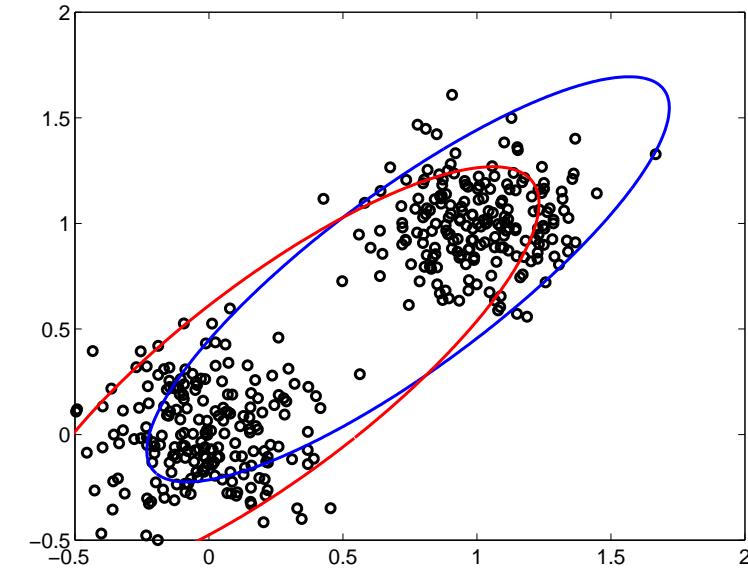
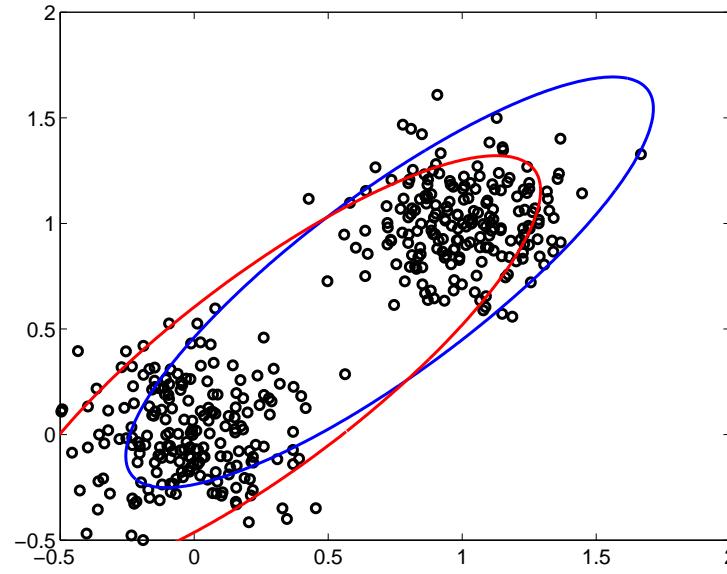
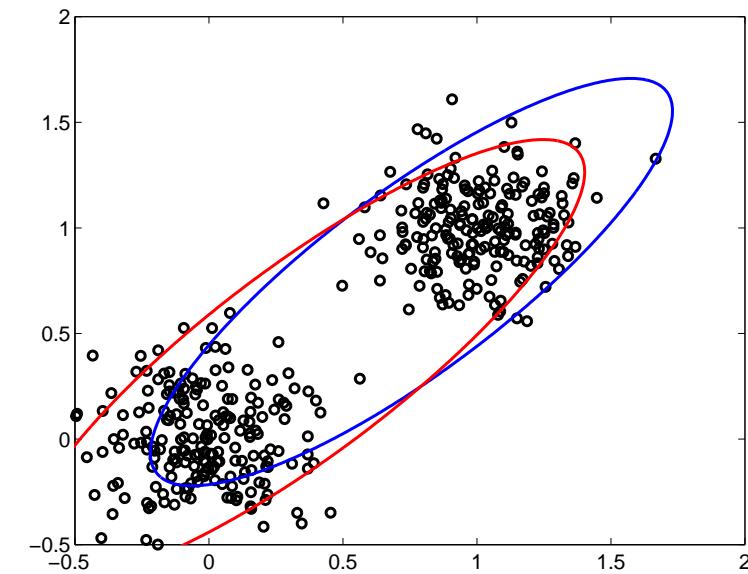
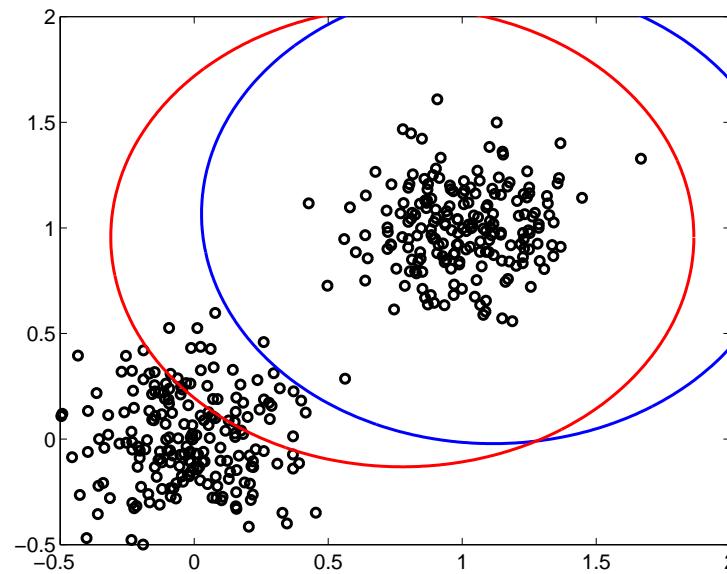
**M-step:** re-estimate the parameters (separately for the two Gaussians) based on the soft assignments.

$$\hat{p}_j \leftarrow \frac{\hat{n}_j}{n}, \text{ where } \hat{n}_j = \sum_{i=1}^n \hat{p}(j|i)$$

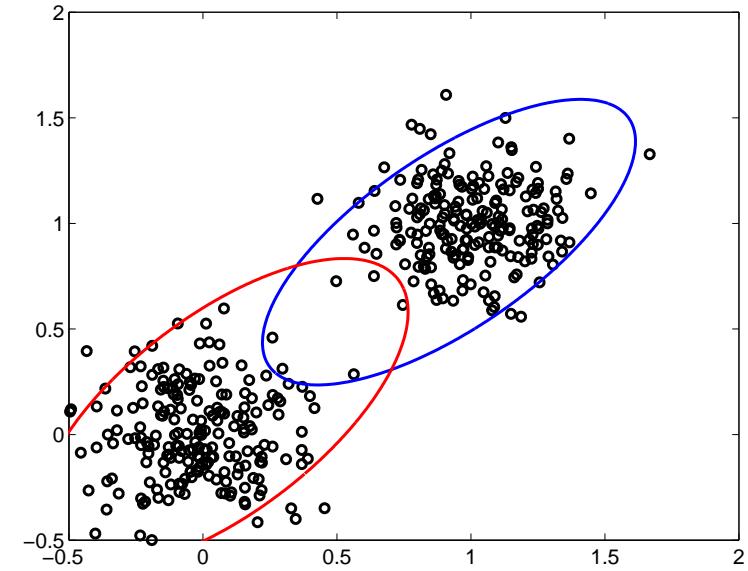
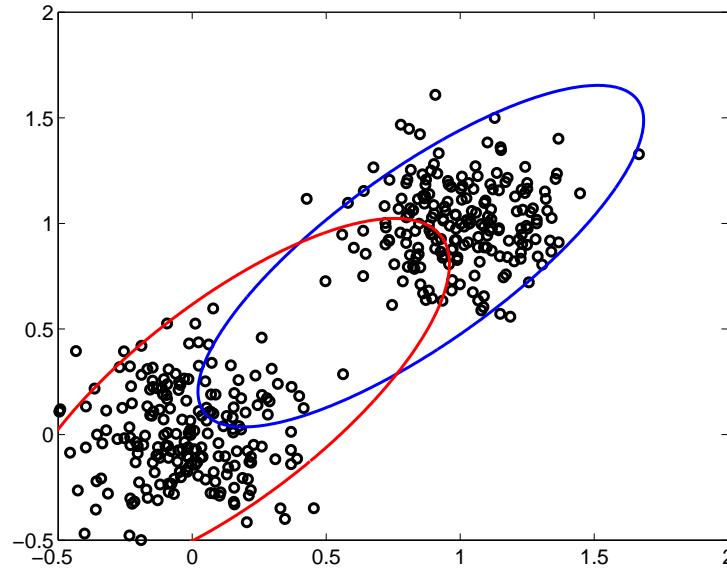
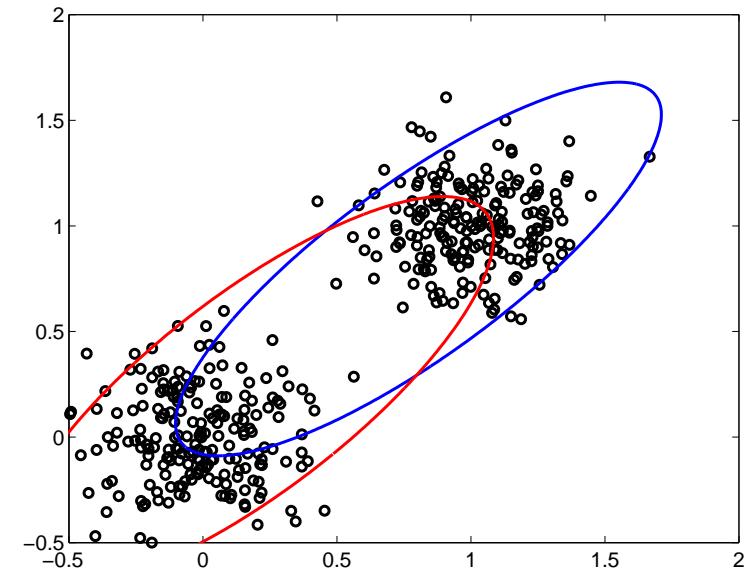
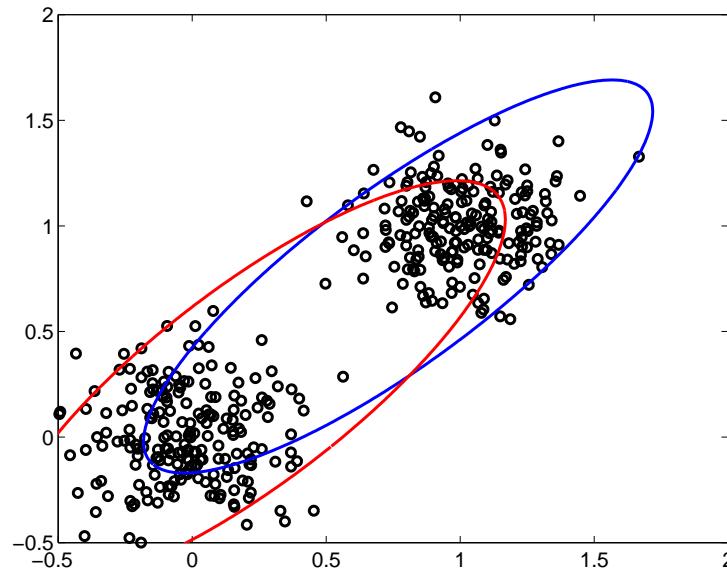
$$\hat{\mu}_j \leftarrow \frac{1}{\hat{n}_j} \sum_{i=1}^n \hat{p}(j|i) \mathbf{x}_i$$

$$\hat{\Sigma}_j \leftarrow \frac{1}{\hat{n}_j} \sum_{i=1}^n \hat{p}(j|i) (\mathbf{x}_i - \hat{\mu}_j)(\mathbf{x}_i - \hat{\mu}_j)^T$$

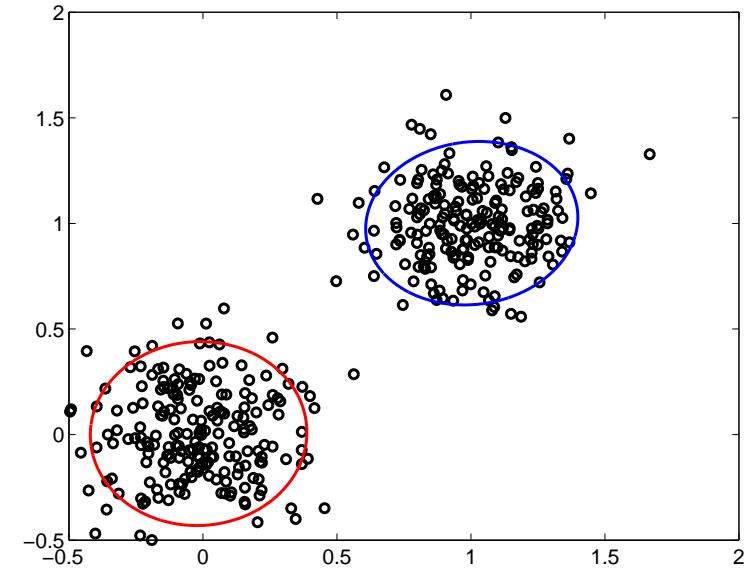
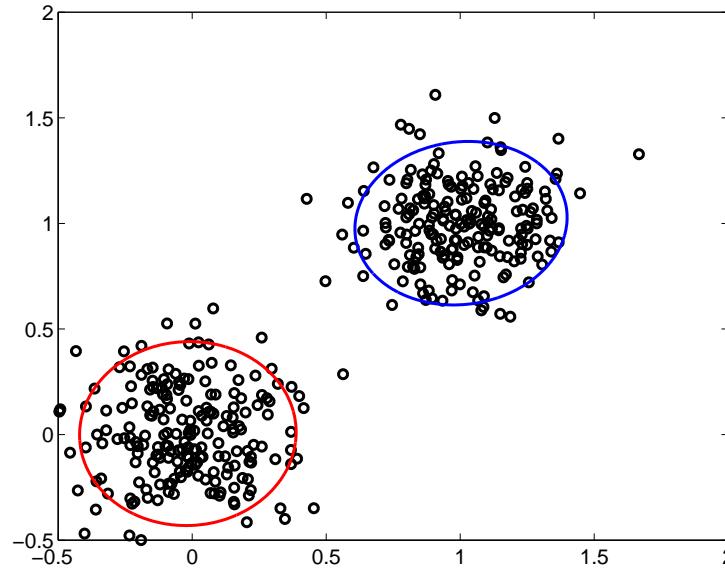
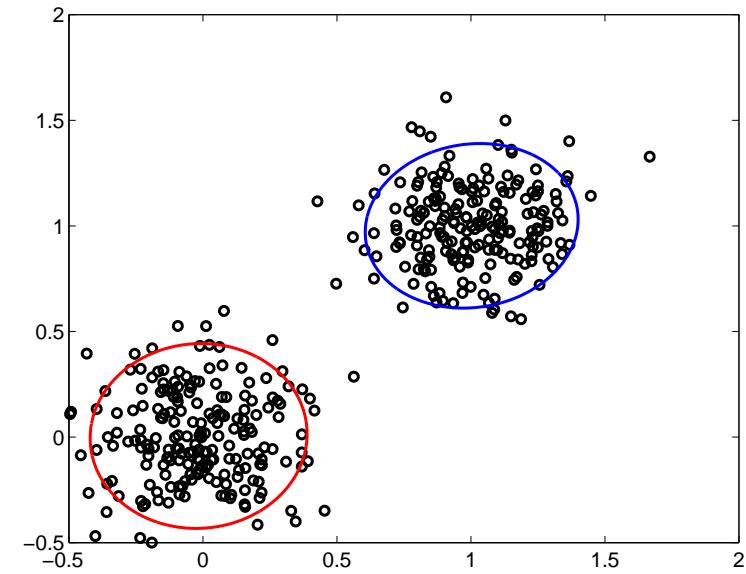
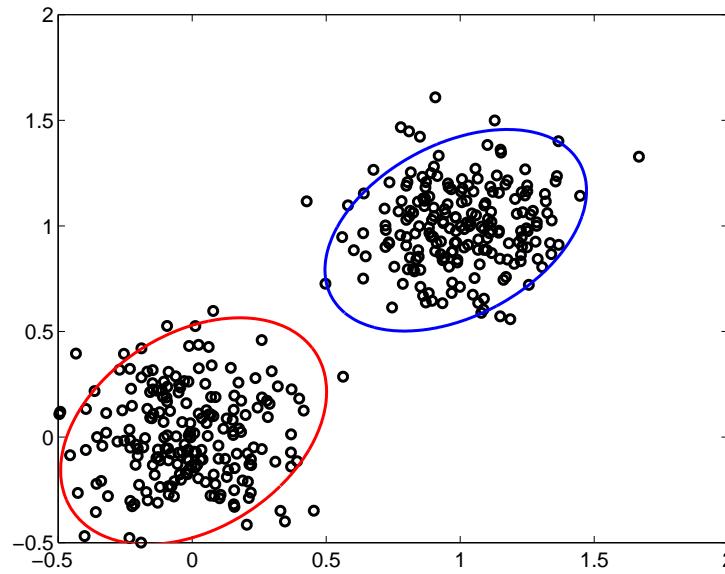
# Mixture density estimation: example



# Mixture density estimation



# Mixture density estimation



# The EM-algorithm

- Each iteration of the EM-algorithm *monotonically* increases the (log-)likelihood of the  $n$  training examples  $\mathbf{x}_1, \dots, \mathbf{x}_n$ :

$$\log p(\text{data} | \theta) = \sum_{i=1}^n \log \left( \underbrace{p_1 p(\mathbf{x}_i | \mu_1, \Sigma_1)}_{p(\mathbf{x}_i | \theta)} + p_2 p(\mathbf{x}_i | \mu_2, \Sigma_2) \right)$$

where  $\theta = \{p_1, p_2, \mu_1, \mu_2, \Sigma_1, \Sigma_2\}$  contains all the parameters of the mixture model.

