



Machine learning: lecture 5

Tommi S. Jaakkola
MIT CSAIL
tommi@csail.mit.edu



Topics

- Classification and regression
 - regression approach to classification
 - Fisher linear discriminant
 - elementary decision theory
- Logistic regression
 - model, rationale
 - estimation, stochastic gradient
 - additive extension
 - generalization



Classification

Example: digit recognition (8x8 binary digits)

binary digit	actual label	target label in learning
	"2"	1
	"2"	1
	"1"	0
	"1"	0
...



Classification via regression

- Suppose we ignore the fact that the target output y is binary (e.g., 0/1) rather than a continuous variable
- So we will estimate a linear regression function

$$f(\mathbf{x}; \mathbf{w}) = w_0 + w_1x_1 + \dots + w_dx_d \\ = w_0 + \mathbf{x}^T \mathbf{w}_1,$$

based on the available data as before.

- Assuming $y = f(\mathbf{x}; \mathbf{w}) + \epsilon$, $\epsilon \sim N(0, \sigma^2)$, then the ML objective for the parameters \mathbf{w} reduces to least squares fitting:

$$J_n(\mathbf{w}) = \frac{1}{n} \sum_{i=1}^n (y_i - f(\mathbf{x}_i; \mathbf{w}))^2$$



Classification via regression cont'd

- We can use the resulting regression function

$$f(\mathbf{x}; \hat{\mathbf{w}}) = w_0 + \mathbf{x}^T \hat{\mathbf{w}}_1,$$

to classify any new (test) example \mathbf{x} according to

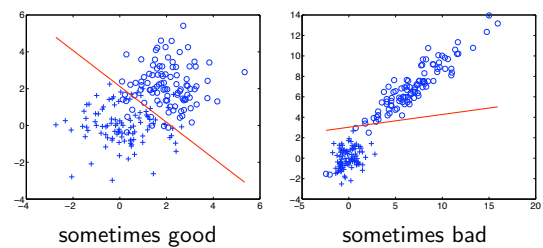
label = 1 if $f(\mathbf{x}; \hat{\mathbf{w}}) > 0.5$, and label = 0 otherwise

- $f(\mathbf{x}; \hat{\mathbf{w}}) = 0.5$ therefore defines a linear *decision boundary* that partitions the input space into two class specific regions (half spaces)



Classification via regression cont'd

- Given the dissociation between the objective (classification) and the estimation criterion (regression) it is not clear that this approach leads to sensible results



Linear regression and projections

- A linear regression function (here in 2D)

$$f(\mathbf{x}; \mathbf{w}) = w_0 + \mathbf{x}^T \mathbf{w}_1$$

projects each point $\mathbf{x} = [x_1 \ x_2]^T$ to a line parallel to \mathbf{w}_1 .

point in \mathcal{R}^d projected point in \mathcal{R}

$$\mathbf{x}_1 \quad z_1 = \mathbf{x}_1^T \mathbf{w}_1$$

$$\mathbf{x}_2 \quad z_2 = \mathbf{x}_2^T \mathbf{w}_1$$

$$\dots \quad \dots$$

$$\mathbf{x}_n \quad z_n = \mathbf{x}_n^T \mathbf{w}_1$$

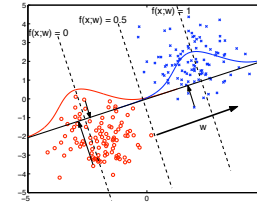
- We can study how well the projected points $\{z_1, \dots, z_n\}$, viewed as functions of \mathbf{w}_1 , are separated across the classes.

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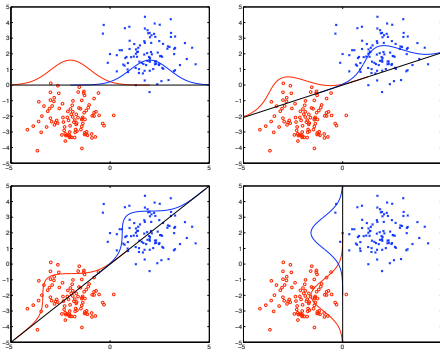
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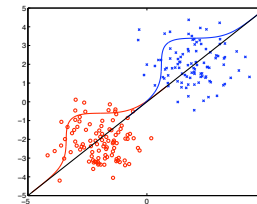
Projection and classification

- By varying \mathbf{w}_1 we get different levels of separation between the projected points



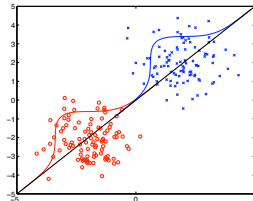
Optimizing the projection

- We would like to find \mathbf{w}_1 that somehow maximizes the separation of the projected points across classes



- We can quantify the separation (overlap) in terms of means and variances of the resulting 1-dimensional class distributions

Fisher linear discriminant: preliminaries



- Class descriptions in \mathcal{R}^d :
 - class 0: n_0 samples, mean μ_0 , covariance Σ_0
 - class 1: n_1 samples, mean μ_1 , covariance Σ_1
- Projected class descriptions in \mathcal{R} :
 - class 0: n_0 samples, mean $\mu_0^T \mathbf{w}_1$, variance $\mathbf{w}_1^T \Sigma_0 \mathbf{w}_1$
 - class 1: n_1 samples, mean $\mu_1^T \mathbf{w}_1$, variance $\mathbf{w}_1^T \Sigma_1 \mathbf{w}_1$

Fisher linear discriminant

- Estimation criterion: we find \mathbf{w}_1 that maximizes

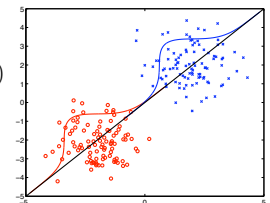
$$J_{Fisher}(\mathbf{w}) = \frac{(\text{Separation of projected means})^2}{\text{Sum of within class variances}}$$

$$= \frac{(\mu_1^T \mathbf{w}_1 - \mu_0^T \mathbf{w}_1)^2}{n_1 \mathbf{w}_1^T \Sigma_1 \mathbf{w}_1 + n_0 \mathbf{w}_1^T \Sigma_0 \mathbf{w}_1}$$

- The solution (class separation)

$$\hat{\mathbf{w}}_1 \propto (n_1 \Sigma_1 + n_0 \Sigma_0)^{-1} (\mu_1 - \mu_0)$$

is decision theoretically optimal for two normal populations with equal covariances ($\Sigma_1 = \Sigma_0$)

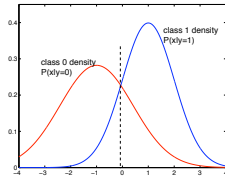




Background: simple decision theory

- Suppose we know the class-conditional densities $p(\mathbf{x}|y)$ for $y = 0, 1$ as well as the overall class frequencies $P(y)$.

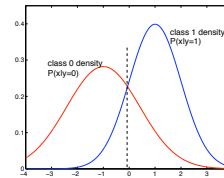
How do we decide which class a new example \mathbf{x}' belongs to so as to minimize the overall probability of error?



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How do we decide which class a new example \mathbf{x}' belongs to so as to minimize the overall probability of error?



The minimum probability of error decisions are given by

$$y' = \arg \max_{y=0,1} \{ p(\mathbf{x}'|y)P(y) \}$$

$$= \arg \max_{y=0,1} \{ P(y|\mathbf{x}') \}$$



Logistic regression

- The optimal decisions are based on the posterior class probabilities $P(y|\mathbf{x})$. For binary classification problems, we can write these decisions as

$$y = 1 \text{ if } \log \frac{P(y = 1|\mathbf{x})}{P(y = 0|\mathbf{x})} > 0$$

and $y = 0$ otherwise.



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$$y = 1 \text{ if } \log \frac{P(y = 1|\mathbf{x})}{P(y = 0|\mathbf{x})} > 0$$

and $y = 0$ otherwise.

- We generally don't know $P(y|\mathbf{x})$ but we can parameterize the possible decisions according to

$$\log \frac{P(y = 1|\mathbf{x})}{P(y = 0|\mathbf{x})} = f(\mathbf{x}; \mathbf{w}) = w_0 + \mathbf{x}^T \mathbf{w}_1$$



Logistic regression cont'd

- Our log-odds model

$$\log \frac{P(y = 1|\mathbf{x})}{P(y = 0|\mathbf{x})} = w_0 + \mathbf{x}^T \mathbf{w}_1$$

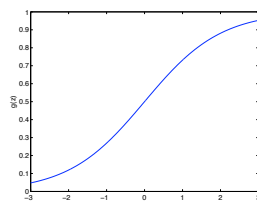
gives rise to a specific form for the conditional probability over the labels (the logistic model):

$$P(y = 1|\mathbf{x}, \mathbf{w}) = g(w_0 + \mathbf{x}^T \mathbf{w}_1)$$

where

$$g(z) = (1 + \exp(-z))^{-1}$$

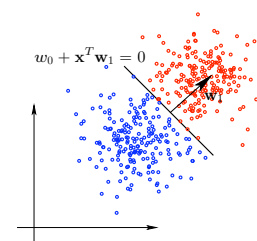
is a logistic "squashing function" that turns linear predictions into probabilities



Logistic regression: decisions

- Logistic regression models imply a linear decision boundary

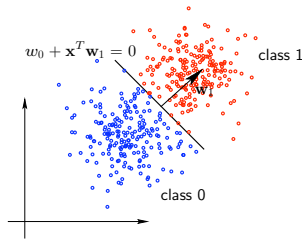
$$\log \frac{P(y = 1|\mathbf{x})}{P(y = 0|\mathbf{x})} = w_0 + \mathbf{x}^T \mathbf{w}_1 = 0$$



Logistic regression: decisions

- Logistic regression models imply a linear decision boundary

$$\log \frac{P(y=1|\mathbf{x})}{P(y=0|\mathbf{x})} = w_0 + \mathbf{x}^T \mathbf{w}_1 = 0$$



Fitting logistic regression models

- As with the linear regression models we can fit the logistic models using the maximum (conditional) log-likelihood criterion

$$l(D; \mathbf{w}) = \sum_{i=1}^n \log P(y_i | \mathbf{x}_i, \mathbf{w})$$

where

$$P(y=1|\mathbf{x}, \mathbf{w}) = g(w_0 + \mathbf{x}^T \mathbf{w}_1)$$

- The log-likelihood function $l(D; \mathbf{w})$ is a *jointly concave* function of the parameters \mathbf{w} ; a number of optimization techniques are available for finding the maximizing parameters

About the ML solution

- If we set the derivatives of the log-likelihood with respect to the parameters to zero

$$\frac{\partial}{\partial w_0} l(D; \mathbf{w}) = \sum_{i=1}^n (y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w})) = 0$$

$$\frac{\partial}{\partial w_j} l(D; \mathbf{w}) = \sum_{i=1}^n (y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w})) x_{ij} = 0$$

the optimality conditions again require that the prediction errors

$$\epsilon_i = (y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w})), \quad i = 1, \dots, n$$

corresponding to the optimal setting of the parameters are uncorrelated with any linear function of the inputs.

Stochastic gradient ascent

- We can try to maximize the log-likelihood in an *on-line* or incremental fashion.

Given each training input \mathbf{x}_i and the binary (0/1) label y_i , we can change the parameters \mathbf{w} slightly to increase the corresponding log-probability

$$\begin{aligned} \mathbf{w} &\leftarrow \mathbf{w} + \eta \frac{\partial}{\partial \mathbf{w}} \log P(y_i | \mathbf{x}_i, \mathbf{w}) \\ &= \mathbf{w} + \eta \underbrace{(y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w}))}_{\text{prediction error}} \begin{bmatrix} 1 \\ \mathbf{x}_i \end{bmatrix} \end{aligned}$$

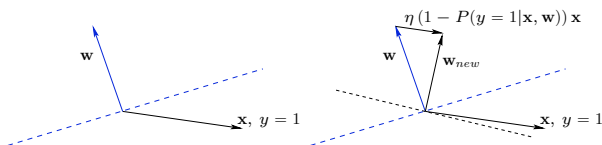
where η is the *learning rate*.

- The resulting update is similar to the mistake driven algorithm discussed earlier; examples that are already confidently classified do not lead to any significant updates

Stochastic gradient ascent cont'd

- To understand the procedure graphically we focus on a single example and omit the bias term w_0

$$\mathbf{w} \leftarrow \mathbf{w} + \eta \underbrace{(y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w}))}_{\text{prediction error}} \mathbf{x}_i$$



Gradient ascent of the log-likelihood

- We can also perform gradient ascent steps on the log-likelihood of all the training labels given examples at the same time. In other words,

$$\begin{aligned} \mathbf{w} &\leftarrow \mathbf{w} + \eta \frac{\partial}{\partial \mathbf{w}} l(D; \mathbf{w}) \\ &= \mathbf{w} + \eta \sum_{i=1}^n (y_i - P(y_i = 1 | \mathbf{x}_i, \mathbf{w})) \begin{bmatrix} 1 \\ \mathbf{x}_i \end{bmatrix} \end{aligned}$$

Still need to figure out a way to set the learning rate to guarantee convergence.



Setting the learning rate: Armijo rule

The learning rate in

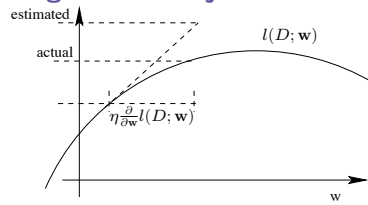
$$\mathbf{w} \leftarrow \mathbf{w} + \eta \frac{\partial}{\partial \mathbf{w}} l(D; \mathbf{w})$$

“should” satisfy

$$l\left(D; \mathbf{w} + \overbrace{\eta \frac{\partial}{\partial \mathbf{w}} l(D; \mathbf{w})}^{\mathbf{w}_{new}}\right) - l(D; \mathbf{w}) \geq \eta \cdot \frac{1}{2} \left\| \frac{\partial}{\partial \mathbf{w}} l(D; \mathbf{w}) \right\|^2$$

The Armijo rule suggests finding the smallest integer m such that $\eta = \eta_0 q^m$, $q < 1$ is a valid choice in this sense.

- Armijo rule is guaranteed to converge to a (local) maximum under certain technical assumptions



Additive models and classification

- Similarly to linear regression models, we can extend the logistic regression models to additive (logistic) models

$$P(y = 1 | \mathbf{x}, \mathbf{w}) = g(w_0 + w_1 \phi_1(\mathbf{x}) + \dots + w_m \phi_m(\mathbf{x}))$$

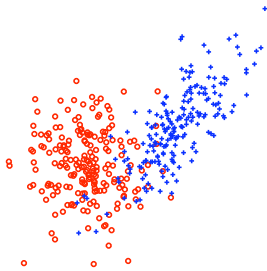
- As before we are free to choose the basis functions $\phi_i(\mathbf{x})$ to capture relevant properties of any specific classification problem
- Since we also over-fit easily, we can use leave-one-out cross-validation (in terms of log-likelihood or classification error) to estimate the generalization performance

$$\text{CV log-likelihood} = \frac{1}{n} \sum_{i=1}^n \log P(y_i | \mathbf{x}_i, \hat{\mathbf{w}}^{-i})$$



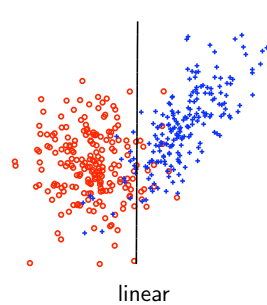
Logistic regression example

- Simple binary classification problem in \mathcal{R}^2



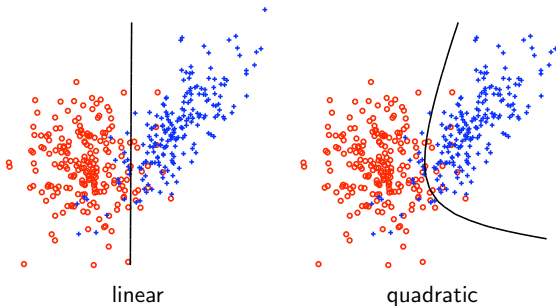
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